

Broadcasting in Sensor Networks: The Role of Local Information

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Abstract—Flooding based strategies are conventionally employed to perform querying and broadcasting in sensor networks. These schemes have low hop-delays of $\Theta(\frac{1}{M(n)})$ to reach any node that is a unit distance away, where $M(n)$ is the transmission range of any sensor node. However, in sensor networks with large radio ranges, flooding based broadcasting schemes cause many redundant transmissions leading to a broadcast storm problem. Many approaches have been proposed to mitigate this problem by utilizing broadcast schemes that employ some knowledge of the previous transmissions to reduce the extraneous transmissions. In this paper, we study the role of geographic information and state information (i.e. memory of previous messages or transmissions) in reducing the redundant transmissions in the network.

We consider three broadcasting schemes with varying levels of local information where nodes have: (i) no geographic or state information, (ii) coarse geographic information about the origin of the broadcast, and (iii) no geographic information, but remember previously received messages. We also consider the related problem of broadcasting to a set of “spatially uniform” points in the network (lattice points) in the regime where all nodes have only a local sense of direction. For each of these networks, we compute the number of transmissions required to achieve broadcast delays that are order-wise equivalent to simple flooding algorithms, i.e. $\Theta(\frac{1}{M(n)})$.

We first show that networks with no geographic or state information require exponentially large number of transmissions whereas networks with very little geographic or state information can utilize the knowledge to significantly reduce the transmission overheads. Next, we show that networks with local information, can reduce the congestion by spreading the messages more uniformly through the network. Finally, we show that networks with only state information can also employ the information to provide a radial drift to the transmitted packets. In the context of lattice broadcasting, we again show that local information results in significant reduction of transmission overheads. We quantitatively compare the transmission overheads of broadcasting strategies and validate our results using simulations.

I. INTRODUCTION

Advances in Micro-embedded computing systems, coupled with developments in wireless technology have enabled the mass production of small sensing devices equipped with wireless communication capabilities. It is envisaged that in the near future sensors networks formed by large-scale deployment of such devices would perform distributed sensing/control

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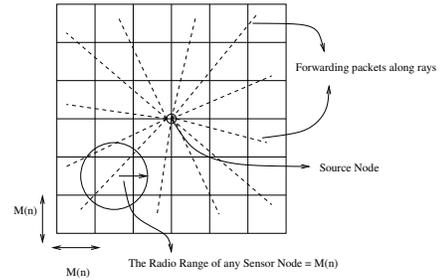


Fig. 1. Forwarding packets along straight lines - We require only one transmission per tile for broadcasting.

operations. Applications for sensor networks include commercial applications involving macro-scale measurements and control, intrusion detection, and robust communication. These networks are characterized by the absence of any established architecture and by constrained energy and computational resources at each node. Communication between any two nodes in these networks is mainly accomplished through packet forwarding by intermediary relay nodes, where messages are relayed to neighbor nodes within the radio range.

In many sensor network applications, broadcasting is a common communication primitive required for various control operations. Applications regularly require broadcast operations to update global information and also to perform network maintenance such as updating topology, route discovery and propagating alarm signals. Similarly, many sensing applications need to periodically inform the sensors to collect information. Thus, an important communication task of a sensor network is to disseminate messages/instructions information to most nodes. A related broadcasting problem arises when a node (say, a controller or a fusion center) needs to query/send a control message to a subset of nodes which are approximately spatially uniform. Such a scenario can arise for instance when the controller needs a spatially uniform sample of a physical underlying process.

In the presence of energy and computation constrained nodes, we require that the communication operations for both these scenarios be energy efficient, computationally simple and delay sensitive. Since the channel utilized by the sensor nodes is a wireless channel, the messages are broadcast to all nodes within the radio range of the transmitting node. Efficient broadcast strategies utilize the inherent broadcast nature of the communication channel to minimize the total number of transmissions, while guaranteeing the reception of the message at all nodes.

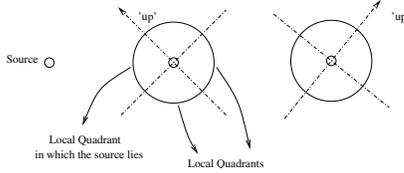


Fig. 2. Local Quadrants in Sensor nodes

Broadcasting in wired networks, is conventionally performed by a simple flooding algorithm, in which each node forwards the message/query once to all its neighbors, when it first hears the message, and ignores all further receptions of the same message. Thus the nodes remember ‘state’ information, i.e, if a nodes has previously heard a message or not. An advantage with such a flooding based broadcast strategy is that it achieves low broadcast delays, without any geographic information at sensor nodes. However, in densely connected networks, such as sensor networks with relatively large radio ranges, such simple flooding based querying/broadcast schemes create many redundant transmissions causing energy inefficiency. Such algorithms lead to a broadcast storm [14] problem, where the same message is received at a node, multiple number of times.

If all nodes in the network had perfect geographic information, it would be possible to considerably reduce the total number of transmissions. Ideally one could use only $(\frac{1}{M(n)})^2$ transmissions in the network, where $M(n)$ is the radio range (in other words, one transmission per tile, see Figure 1). Whereas with simple flooding, the number of transmissions would scale as n (the number of nodes), which could be much larger in dense sensor networks.

A simple scheme to achieve a $\Theta(\frac{1}{M(n)})^2$ number of transmissions¹ is by dispatching packets along “rays” as shown in Figure 1. This scheme requires “perfect” geographic information at the nodes in order to make sure that the rays do not “bend” or loop back.

Similarly, under perfect state information, where all nodes had knowledge of past transmissions and routing tables, it is possible to reduce the broadcast redundancy by constructing a minimum spanning tree or creating an overlay network. However, in many sensor networks, it is impractical to acquire perfect geographic information, as it requires sophisticated location devices and/or computational capabilities. In networks with simultaneous broadcasts by many sources, nodes are required to maintain routing information for messages from each source. Thus it is infeasible to store all routing state information, in networks with meager storage resources.

In this paper, we study the role of information (geographic and state information) on reducing the broadcast redundancy, while preserving the delay efficiency of flooding based approaches for the cases of (i) broadcasting over the entire space, and (ii) broadcasting over a lattice. In particular, we consider efficient broadcast strategies in networks with varying levels

¹By $g(n) = O(f(n))$, we note that there exists a positive constant c_1 such that for $n > N$, $g(n) \leq c_1 f(n)$. We say $g(n) = \Omega(f(n))$, if $f(n) = O(g(n))$. $g(n) = \Theta(f(n))$ if $g(n) = O(f(n))$ and $f(n) = O(g(n))$.

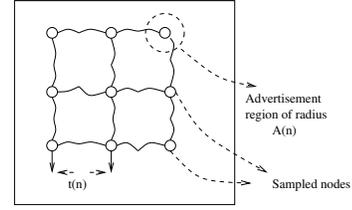


Fig. 3. Spatial sampling in Sensor nodes

of information at the nodes:

- 1) *Zero information*: Nodes have no geographic information or memory, and broadcasting is only through random packet forwarding.
- 2) *Source Quadrant Information*: Nodes have a local notion of four directions, not common to all nodes, and know the local quadrant in which the source is located (See Figure 2).
- 3) *Transmission state Information*: Nodes have no geographic information, but all nodes remember messages received previously. (i.e. state information)
- 4) *Local Direction Information*: Finally, for the problem of spatial sampling (see Figure 3), we consider nodes that have an approximate sense of ‘East’, ‘West’, ‘North’ and ‘South’, but have no other geographic information, source or receiver location information or memory.

A. Main Contributions

We consider dense sensor networks on a plane, where each node has a large number of neighbor nodes, within its radio range $M(n)$ ². We model this by a continuum of sensor nodes, where we associate a sensor node to every point in the plane. We measure broadcast delay in terms of the number of hops required to reach any given point on the network, ignoring the queuing delay at the nodes (for a similar model, see [21],[11]). Under this network model, we quantitatively analyze the efficiency of broadcast strategies with varying levels of information at the sensor nodes.

We first observe that flooding-based strategies lead to broadcast delays that are of the same order as optimal straight-line broadcasting (although with many more transmissions), i.e., the broadcast delay $D(n) = \Theta(\frac{1}{M(n)})$. Thus, in order to compare broadcasting in networks with varying levels of information, we restrict the strategies to possess a broadcast delay $D(n) = \Theta(\frac{1}{M(n)})$, i.e., are order-wise equivalent to flooding based strategies. For a broadcast to reach a node, it is necessary for a transmission to occur within the radio range $M(n)$ of the sensor node. However, if all the neighbor nodes within a radius $A(n)$ contain routing information to direct the transmission to the intended node, the transmissions are only required to reach a ball of radius (advertisement) $A(n)$ about the node.

We measure delay in terms of hop-count, and energy efficiency in terms of total transmissions per search/broadcast and quantitatively analyze the information vs. efficiency trade-off

²The parameter n roughly corresponds to the density of the network.

Information type	Transmissions $T(n)$	Congestion
Zero Information	$c^{\frac{1}{M(n)}}, c > 1$	Heavy congestion about the source
Source Quadrant Information	$(\frac{1}{M(n)})^2$	Moderate congestion about the source
Transmission State Information	$(\frac{1}{M(n)})^2 \log \frac{1}{M(n)}$	Low congestion throughout the network

TABLE I
TRADE-OFFS IN BROADCASTING - NETWORKS WITH LIMITED INFORMATION

in networks. The trade-offs are provided in Table I. We show the following results on broadcast efficiency for networks with varying levels of information.

- (i) In networks with zero information, we present broadcast strategies based on random packet relaying. We show that an exponentially large number of transmissions (of the order of $c^{\frac{1}{M(n)}}$, for some $c > 1$) is necessary to ensure a transmission within a radius $M(n)$ of any given node at a unit distance from the source of the broadcast, to achieve a broadcast delay of $\Theta(\frac{1}{M(n)})$. Further, we show that exponentially large number of transmissions are sufficient for achieving this broadcast delay. To demonstrate this, we employ an inequality result on the concentration of the probability measures for sums of i.i.d random variables. We also show that there are a large number of simultaneous transmissions in the region surrounding the source node, thus causing congestion in that area.
- (ii) We consider networks with source quadrant information, and present a broadcast strategy based on packet forwarding that provides radial drift to the transmissions. We show that the outward spread of the transmissions reduce the broadcast redundancy. We show that only $(\frac{1}{M(n)})^2$ transmissions are required to achieve a delay $D(n) = \Theta(\frac{1}{M(n)})$.
- (iii) In networks with state information, we show that broadcast strategies can learn to inherently provide a radial drift to the transmissions. The broadcast strategies can use the state information to suppress transmissions by redundant nodes and advance the packets away from the source of the broadcast. By considering a strategy of suppression based on [9], we show that this implied radial drift suffices to achieve optimal broadcast delays with $T(n) = (\frac{1}{M(n)})^2 \log 1/M(n)$, and negligible congestion throughout the network.
- (iv) For the problem of spatial sampling in networks with local direction information, we present a randomized-tree based broadcast strategy that provides a lower transmission overhead. We show that we can sample on a “grid” of a given size $s(n) \sim \Omega(\frac{1}{\log \frac{1}{M(n)}})$ (see Figure 3), as long as the “bin” size (the local advertisement radius) scales as $(M(n))^\gamma$, $\gamma < .49$. Such a sampling requires the number of transmissions to scale as $T(n) = (\frac{1}{M(n)})^\alpha$, for a finite α that depends on $s(n)$.

Finally, we provide continuum model based simulations which support the analytical results obtained in the paper. From the above results we infer that broadcasting with very little geographic or state information is significantly more efficient than networks without any such “local” knowledge.

While strategies with local state information can provide low transmission overheads and low congestion, the memory requirement scales linearly with the number of simultaneous broadcast messages. Further, such strategies also require nodes to compare the messages in their memory with every received broadcast. On the other hand, we note that strategies with geographic location information also provide substantial reductions in the number of transmissions, and the information requirements do not increase with simultaneous broadcasts. However, obtaining geographic information at the sensor nodes might require significant computation and/or hardware, such as GPS. In practice, these considerations can be used to trade-off between memory, hardware and energy efficiency (number of transmissions).

B. Related Work

There has been considerable work on broadcasting and querying in sensor networks [23], [19], [2], [20], [14], [16], [4], [1], [13], [12], [17]. It was demonstrated in [14] that flooding based broadcasting/querying schemes such as [12], [13] cause many redundant transmissions leading to the broadcast storm problem. As discussed in [24], many of the broadcast schemes introduced to mitigate the “Broadcast storm” problem can be classified into the following categories:

- (i) Probabilistic schemes such as [9], [15] in which nodes that receive the message rebroadcast with a fixed probability. In these schemes the nodes are assumed to have state information to remember previously received messages and utilize them to suppress secondary transmissions.
- (ii) Location based schemes proposed in [14] where node transmission decisions are based on the expected area covered.
- (iii) Neighbor knowledge based methods such as [2], [16], [20] where the location of the neighbors or the two-hop neighbors are known. In [23], perfect information about the position of all nodes in the network is utilized to construct minimum energy broadcast trees, whereas in [2], the authors provide a construction for a similar tree based on local topology information. In [20], the two-hop neighbor information is utilized for building connected dominating sets that efficiently broadcast information.

Also, querying in sensor networks has been studied in papers such as [17], [5]. The authors in [1] propose random walks initiated by the source node and the destination node. They have shown that querying delay, transmission overheads can be

reduced by spreading routing information through the network. This phenomenon has been quantitatively studied in [18].

Although many of the broadcast strategies previously discussed utilize some kind of local knowledge or state information, we note that a systematic analysis of the role of information in broadcasting, and the related trade-offs in the number of transmissions, delay and congestion, has not been explored previously. In this paper, we study a sequence of networks with varying levels of geographic and state information, and compare broadcast trade-offs through analytical methods. Furthermore, we provide simulation results to validate the analytical studies.

II. SYSTEM DESCRIPTION

A. Network Model

We consider a sensor network in which the sensor nodes are deployed over a planar region. Each of these sensor nodes are assumed to have a common circular transmission region and are connected to all other sensor nodes that lie within its transmission radius. The transmission radius is set to scale as $M(n)$, where n is the scaling parameter.³ In this paper, we study broadcast strategies in dense networks in the large- n regime, (where $n \rightarrow \infty$). The results of [8] show that for $M(n) = \Omega\left(\sqrt{\frac{\log n}{n}}\right)$, the network formed by the collection of sensor nodes in a given region of finite area is asymptotically connected, and more importantly the number of nodes in the transmission radius of each node in that given region *tends to infinity* asymptotically. In this paper, we consider any $M(n)$ that scales as $O\left(\frac{1}{n^p}\right)$, $p \in (0, \frac{1}{2})$, to model the growth of the network size relative to the radio range.

Motivated by the above results, in this paper we assume a *continuum model* of the sensor network, where any point in the radio range of a transmitting node S can receive the packet transmitted by the node, and can act as a retransmission node. The neighbor set (nodes within the radio range) of any node S in the sensor network is defined as

$$\mathcal{N}_{S, M(n)} = \{X \in \mathbb{R}^2 : d(X, S) < M(n)\}, \quad (1)$$

where d is the Euclidean distance. Thus, there is a one-to-one correspondence between nodes and their locations and the discretization effects due to node locations are ignored in the continuum model. However, as mentioned above, in densely connected sensor networks, the number of nodes within the radio range of any particular node increases to infinity [8]. Thus, the continuum model appears reasonable in this regime. We refer to [21] for a comparison of analytical results using a continuum model and simulation results with a discrete model with a dense network of nodes, which indicate that the discretization effects are not significant.

B. Broadcast Model

Querying and Information spreading, are both studied as a series of packet forwards in a sensor network. Since the transmissions in a wireless sensor network are inherently

broadcast transmissions, we assume that whenever a node transmits a query, all nodes in its neighbor set can potentially receive it without error.

To broadcast a query ‘‘m’’, the originating node S_0 sends out a packet, to all its neighbors (in a single transmission) and requests a subset \mathcal{S}_1 of its neighbors to retransmit it. The repeated application of this process disseminates the information/query into the network. Let $S_0 = \mathbf{0}$ be the position of the source node, i.e., the position of the node initiating the query. The set \mathcal{S}_i consists of all points (nodes associated with the points) in the network that transmit the packet at the i^{th} iteration of the process, or are the i^{th} generation transmitters.

In this setup, we define the normalized broadcast delay, $D(n)$ as follows. Let X be any given point (or the node at position X), a unit distance away from the origin. We define

$$D(n) = \inf\{i : d(\mathcal{S}_i, X) < M(n)\} \quad (2)$$

where d is the Euclidean distance metric. That is, the normalized delay $D(n)$ is defined as the smallest iteration by which there is a transmission within the radio range of the given point/node X . Note that the ‘unit’ distance between the node and the origin is arbitrary. For any other distance r , the hop-delay can be scaled accordingly. Thus, we define $D(n)$ as the hop count of the minimum hop path from the source to reach any arbitrarily chosen node X that is a unit distance away.

We note that in this definition, the medium access delay has been ignored, and delay is measured only in terms of the hop count. We note that the actual packet delay can be decomposed into the hop count delay and the MAC delay. By suitably scaling the packet size (see [6] for this approach), we can achieve a MAC delay that is order-wise smaller than the hop-count delay. Thus, in this regime, the hop-count will be representative of the packet delay. Even if such a packet-scaling was not employed, the delay with two broadcasting schemes can be compared using a pair of metrics: (i) hop-count delay, and (ii) the ‘‘local’’ congestion about a transmitting node (i.e., the number of transmissions that occur in a spatial region) which clearly plays an important part in determining the MAC delay. Thus, in addition to small hop-count, a good broadcasting scheme will mitigate local congestion. This observation motivates us to later consider ‘‘branching’’ based schemes where the number of transmissions progressively increases with radial distance from the source, and have (order-wise) the same hop-count as more ‘‘concentrated’’ broadcasting schemes.

In the case of querying, the normalized delay corresponds to the number of iterations required to reach any given point that is located a unit distance away from the source node. In the context of information spreading, the normalized delay corresponds to the iterations required to spread the information to a randomly chosen point which is a unit distance away from the source. Thus, the above definition of delay allows us to study the symmetric problems of information spreading and querying within the same framework.

We also define the transmission overhead $T(n)$ as the total number of transmissions by the iteration $D(n)$, i.e. $T(n) = \sum_{k=1}^{D(n)} |\mathcal{S}_k|$, where $|\cdot|$ denotes the cardinality of the set. Conventional flooding based strategies achieve a broadcast

³The quantity n roughly corresponds to the density of nodes in the network.

delay of $\Theta(\frac{1}{M(n)})$ hops in densely connected networks, as the minimum distance of order $M(n)$ is covered in each iteration along all directions. In order to compare the various broadcast strategies, we constrain the broadcast strategies in all network models to achieve order-wise optimal hop-delays. Further, if the delay is $\frac{K}{M(n)}$, $K < \infty$, for any arbitrarily picked node, the broadcast can be efficiently terminated by setting TTL values in the broadcast packets appropriately. Thus, in the rest of the paper, we only consider strategies that have a delay of $\Theta(\frac{1}{M(n)})$.

III. BROADCASTING IN NETWORKS WITH ZERO INFORMATION.

In this section, we study the energy-delay trade-offs of broadcasting, in networks with zero information. We assume that the nodes in the network do not have any geographic or state information. That is, the nodes have no knowledge of the locations of their neighbors or of the broadcast source, and are incapable of remembering previous messages or transmission routes. Since nodes have no state information, decisions to retransmit a received message are made at the time of arrival of the message. Thus, it is possible for the same message to be received and transmitted multiple times by a node.

To broadcast information in such networks with very limited capability, we employ a simple broadcast strategy based on random packet forwarding, that requires no state or geographic information. In this scheme, each transmitting node selects only one retransmitting node randomly from its neighbor set (the nodes within the radio range $M(n)$), and requests the node to retransmit it. We study these “random walk” based schemes, as they are a sequence of simple communication operations and representative of broadcast strategies possible in networks with no information.

As discussed earlier in Section II, to compare the different broadcast schemes, we require the normalized broadcast delay, $D(n) = \Theta(\frac{1}{M(n)})$. By randomly forwarding a single message, it may not be possible to achieve the required normalized delay and hence we initiate multiple broadcasts of the same message, corresponding to independent parallel random walks. That is, we originate $R(n)$ independent copies of the same broadcast message at the source node, and propagate each message by random packet forwarding.

To analyze the energy efficiency of the broadcast strategy, we choose a random node that is a unit distance away from the source node, and compute the total number of transmissions $T(n)$ that are required to ensure that the message is received by the chosen node, within $\Theta(\frac{1}{M(n)})$ iterations. The energy efficiencies are studied in terms of the number of broadcasts.

A. Random Packet Forwarding

The packet forwarding based broadcast, with multiple copies of the broadcast message, has a simple communication structure. The source node transmits $R(n)$ independent copies of the broadcast message, i.e., for every copy of the message, the source node picks another sensor node randomly from its neighbor set for retransmission. Every transmitting node has only one offspring node, and only one transmission per

query/message occurs at every iteration. That is, at the i^{th} iteration, the position of the transmitting node for the k^{th} copy of the message is

$$S_i^k = S_{i-1}^k + X_i^k, \quad (3)$$

where denotes S_l^k the position of the k^{th} random walk after l iterations and X_i^k is the random displacement from the node transmitting copy k at iteration $i - 1$. We assume that X_i^k are i.i.d random variables, with a common distribution μ . Since no geographic location information is available, we assume that the next hop nodes are chosen uniformly randomly from the neighbor set of each transmitter and assume that the distribution μ is uniformly distributed over the compact set $B_{M(n)}(0)$, where $B_r(x)$ denotes a ball of radius r around x . We use the following notation for n -fold convolutions of μ , (i.e., the distribution of n random variables with distribution μ)

$$\mu_{(n+1)}(A) := \int \mu_{(n)}(A - x)\mu(dx), \quad n \in \mathbb{N}, \quad (4)$$

where $\mu_{(1)} := \mu$.

For the above model of a network with Zero Information, we show that the number of transmissions increasing exponentially with $\frac{1}{M(n)}$ (the network diameter in hops), are necessary and sufficient to ensure an optimum broadcast delay. The following theorem shows that exponentially large number of transmissions are necessary to achieve a delay $D(n) = \Theta(\frac{1}{M(n)})$, using the broadcasting strategy discussed earlier in this section. We show that, even if the number of paths are exponentially large, the probability that none of the paths reach the radio range of the node within $\Theta(\frac{1}{M(n)})$ steps is high.

Theorem 3.1: For any given $K < \infty$, there is a $c > 1$ such that for $R(n) = c^{\frac{K}{M(n)}}$,

$$\mathbb{P}\left(\bigcap_{\substack{l=1, \dots, R(n) \\ k=1, \dots, \frac{K}{M(n)}}} S_l^k \notin B_\epsilon(x)\right) \longrightarrow 1.$$

for some $\epsilon > 0$.

Proof:

$$\begin{aligned} & \mathbb{P}\left(\bigcap_{\substack{l=1, \dots, R(n) \\ k=1, \dots, \frac{K}{M(n)}}} S_l^k \notin B_\epsilon(x)\right) \\ &= \mathbb{P}\left(\bigcap_{k=1, \dots, \frac{K}{M(n)}} S_k^1 \notin B_\epsilon(x)\right)^{R(n)} \\ &\geq \left(1 - \sum_{k=1}^{\frac{K}{M(n)}} \mathbb{P}(S_k^1 \in (B_\epsilon(x)))\right)^{R(n)} \\ &\geq \left(1 - \frac{K}{M(n)} \max_k \mathbb{P}(S_k^1 \in (B_\epsilon(x)))\right)^{R(n)}. \end{aligned}$$

Note that $\mathbb{P}(S_k^1 \in B_\epsilon(x)) \leq \mathbb{P}(S_k^1 \in B_{1-\epsilon}^c(0))$. By Chernoff's bound, $\mathbb{P}(S_n^1 \in B_{1-\epsilon}^c(0)) \leq e^{-nI(\delta)}$, for small $\delta > 0$, where I is the rate function associated with the random

variables. Hence,

$$\begin{aligned} & \mathbb{P}\left(\bigcap_{\substack{l=1,\dots,R(n), \\ k=1,\dots,\frac{K}{M(n)}}} S_k^l \notin B_\epsilon(x)\right) \\ & \geq \left(1 - \frac{K}{M(n)} \left(e^{-\frac{K}{M(n)}I(\delta)}\right)\right)^{R(n)}. \end{aligned} \quad (5)$$

Let $R(n) = c \frac{K}{M(n)}$ for any $c < e^{-I(\delta)/2}$ then, the R.H.S term tends to 1 as $n \rightarrow \infty$. ■

Thus, the total number of required transmissions, $T(n) = c \frac{K}{M(n)} \times \frac{2}{M(n)}$, grows exponentially with the network diameter $\frac{1}{M(n)}$.

We now show in following theorem that it is also *sufficient* to have exponentially large number of transmissions, to achieve a delay of $\Theta(\frac{1}{M(n)})$, using the broadcasting strategy discussed earlier in this section.

Theorem 3.2: Consider $R(n)$ independent random walks starting from $S_0 = \mathbf{0}$ and any given point $x = (x_1, x_2)$ on the boundary of the compact ball $B_1(0)$. Then, there exists a $c < \infty$ such that for $R(n) \geq c \frac{1}{M(n)}$,

$$\lim_{n \rightarrow \infty} \min_{k \in 1, \dots, R(n)} \|S_{\frac{2}{M(n)}}^k - x\| - M(n) \leq 0. \quad (a.s)$$

That is, there exists a random walk that is arbitrarily close to x , after $\frac{2}{M(n)}$ iterations.

Proof: We prove the above claim by using Borel-Cantelli's lemma and showing that, the probability that none of the random walks are 'close' to the point x is 'exceedingly' small. That is, to prove that $\min_{k \in 1, \dots, R(n)} \|S_{\frac{2}{M(n)}}^k - x\| \leq M(n)$, we show that

$$\sum_n \mathbb{P}\left(\bigcap_{k \in 1, \dots, R(n)} S_{\frac{2}{M(n)}}^k \notin B_{M(n)}(x)\right) < \infty.$$

$$\begin{aligned} & \mathbb{P}\left(\bigcap_{k \in 1, \dots, R(n)} S_{\frac{2}{M(n)}}^k \notin B_{M(n)}(x)\right) \\ & = \{\mathbb{P}(S_{\frac{2}{M(n)}}^1 \notin B_{M(n)}(x))\}^{R(n)} \end{aligned} \quad (6)$$

$$= \{1 - \mu_{(\frac{2}{M(n)})}(B_{M(n)}(x))\}^{R(n)}, \quad (7)$$

Equation 6 is due to the independence of the random walks, and Equation 7 follows, as $S_{\frac{2}{M(n)}}^1$ is the sum of i.i.d random variables, with distribution μ . To provide an upper bound for $\mathbb{P}(S_{\frac{2}{M(n)}}^1 \notin B_{M(n)}(x))$ (i.e. a lower bound on $\mu_{(\frac{2}{M(n)})}(B_{M(n)}(x))$), we need the following claim.

Claim 1: There exists sub-probability measures ν_1, ν_2 and constants $\delta_1, \delta_2 > 1$ such that

- 1) $\delta_1 \nu_1, \delta_2 \nu_2$ are probability distributions, symmetric about $\frac{M(n)}{2}x_1$ and $\frac{M(n)}{2}x_2$,
- 2) $\sigma_{\delta_1}^2, \sigma_{\delta_2}^2 > 0$, with compact supports $B_1, B_2 \subset \mathbb{R}$ respectively,
- 3) $\mu_n(B_\epsilon(x)) \geq (\delta_1 \delta_2)^{-n} \phi_n(B_\epsilon(x))$, where $\phi = \delta_1 \nu_1 \times \delta_2 \nu_2$, i.e., the product distribution.

We construct a sub-probability measure ν such that $\nu = \mu$ on the set $A \subset B_{M(n)}(0) \subset \mathbb{R}^2$ and zero elsewhere, where

$$A = \{y : \|y - \frac{M(n)}{2}x\|_{L_\infty} < \frac{M(n)}{10}\}.$$

Thus, if (x_1, x_2) are the components of the vector x , the measure ν is defined on the product space

$$\begin{aligned} A = & \left[\frac{M(n)}{2}(x_1 - \frac{1}{5}), \frac{M(n)}{2}(x_1 + \frac{1}{5})\right] \\ & \times \left[\frac{M(n)}{2}(x_2 - \frac{1}{5}), \frac{M(n)}{2}(x_2 + \frac{1}{5})\right]. \end{aligned}$$

Hence, the measure ν over the set A can be expressed in product form as follows:

$$\nu = \nu_1 \times \nu_2,$$

where ν_1 is a measure on $B_1 = [\frac{M(n)}{2}(x_1 - \frac{1}{5}), \frac{M(n)}{2}(x_1 + \frac{1}{5})]$ and ν_2 is a measure on $B_2 = [\frac{M(n)}{2}(x_2 - \frac{1}{5}), \frac{M(n)}{2}(x_2 + \frac{1}{5})]$.

Now, let $\delta_1, \delta_2 > 1$ such that $\delta_1 \nu_1, \delta_2 \nu_2$ are probability distributions over their respective supports. Then, $\phi = \delta_1 \nu_1 \times \delta_2 \nu_2$ is a probability measure on A .

To prove Claim 1 (iii), we establish the following lemma.

Lemma 3.1: Let μ, ν and ϕ be as defined above. Then

$$\mu_{(l)}(x) \geq \nu_{(l)}(x) = \frac{1}{(\delta_1 \delta_2)^l} \phi_{(l)}(x)$$

Proof: We prove this by induction. Clearly, for $l = 1$, we have $\mu(x) \geq \nu(x)$. Let us assume that for $l - 1$, $\mu_{(l-1)}(x) \geq \nu_{(l-1)}(x)$. Since $\mu_{(l)} = \mu_{(l-1)} * \mu$, we have

$$\begin{aligned} \mu_{(l)}(x) & = \int_{-\infty}^{\infty} \mu_{(l-1)}(y) \mu(x - y) dy \\ & \geq \int_{-\infty}^{\infty} \nu_{(l-1)}(y) \nu(x - y) dy \\ & = \nu_{(l)}(x) \end{aligned} \quad (8)$$

$$\Rightarrow \mu_{(l)}(x) \geq \nu_{(l)}(x) = \frac{1}{(\delta_1 \delta_2)^l} \phi_{(l)}(x) \quad (9)$$

We require the following corollary of the result in [7] *Thm. 1, pg. 533*, on the concentration of the distribution about its mean.

Lemma 3.2: We assume that ψ is a probability measure with mean t , variance $\sigma_\psi^2 > 0$, and a compact support $B \subset \mathbb{R}$. Then, for some $K < \infty$,

$$\psi_n\left(\left[nt - \frac{\epsilon}{2}, nt + \frac{\epsilon}{2}\right]\right) \geq \frac{K}{\sqrt{n}}, \quad \forall n > N_0 \in \mathbb{N}.$$

In our scenario, where the support for the distribution ϕ is not unit length but within a $c_1 M(n)$ length interval, the result can be modified to show that

$$\begin{aligned} & \phi_{\frac{2}{M(n)}}\left(\left[x_1 - \epsilon' M(n), x_1 + \epsilon' M(n)\right]\right) \\ & \times \left[x_2 - \epsilon' M(n), x_2 + \epsilon' M(n)\right] \geq \frac{KM(n)}{2}. \end{aligned} \quad (10)$$

To show Theorem 3.2, we apply Lemma 3.1 to Equation 7, to see that

$$\begin{aligned} \mathbb{P}(S_{\frac{2}{M(n)}}^1 \notin B_{M(n)}(x)) & = 1 - \mu_{(\frac{2}{M(n)})}(B_{M(n)}(x)) \\ & \leq 1 - (\delta_1 \delta_2)^{-\frac{2}{M(n)}} \phi_{(\frac{2}{M(n)})}(B_{M(n)}(x)). \end{aligned} \quad (11)$$

Notice that

$$\begin{aligned} & \phi_{(l)}(B_{M(n)}(x)) \\ & \geq \delta_1 \nu_{1(l)}\left[l \frac{M(n)}{2} x_1 - \epsilon' M(n), l \frac{M(n)}{2} x_1 + \epsilon' M(n)\right] \\ & \times \delta_2 \nu_{2(l)}\left[l \frac{M(n)}{2} x_2 - \epsilon' M(n), l \frac{M(n)}{2} x_2 + \epsilon' M(n)\right]. \end{aligned} \quad (12)$$

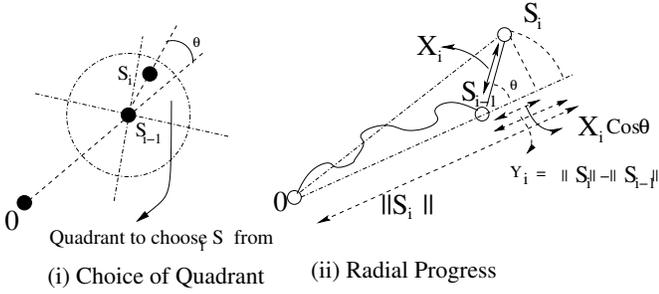


Fig. 4. Random packet forwarding with knowledge of source location, and local quadrants

Applying Lemma 3.2 to the distributions $\delta_1\nu_1$ and $\delta_2\nu_2$, we see that for a large enough K ,

$$\phi_{(l)}(B_{M(n)}(x)) \geq \frac{K}{l} \quad \forall l \in \mathbb{N}. \quad (13)$$

It follows by Equation 11 that

$$\begin{aligned} & \{\mathbb{P}(S_{\frac{2}{M(n)}}^1 \notin B_{M(n)}(x))\}^{R(n)} \\ &= \{1 - \mu_{(\frac{2}{M(n)})}(B_{M(n)}(x))\}^{R(n)}, \end{aligned} \quad (14)$$

$$\leq \left\{1 - (\delta_1\delta_2)^{-\frac{2}{M(n)}} \frac{KM(n)}{2}\right\}^{R(n)}, \quad (15)$$

Let $c > (\delta_1\delta_2)^2$. Then, it is seen that

$$\sum \left\{1 - (\delta_1\delta_2)^{-\frac{2}{M(n)}} \frac{KM(n)}{2}\right\}^{c^{\frac{1}{M(n)}}} < \infty \quad (16)$$

Thus, for $R(n) = c^{\frac{1}{M(n)}}$, we show, by using Borel-Cantelli's lemma that

$$\lim_{n \rightarrow \infty} \min_{k \in \{1, \dots, R(n)\}} \|S_{\frac{2}{M(n)}}^k - x\| - M(n) \leq 0. \quad (a.s.)$$

Remark 3.1: Thus, the results in this section indicate that exponentially large number of transmissions are necessary and sufficient to successfully broadcast in sensor networks with no geographic or state information. We also note that, by employing multiple queries/messages, the number of transmissions by nodes close to the source node increases linearly with the $R(n)$. In networks with Zero Information, this translates to an exponentially large number of transmissions in a small area (areas of the size of the radio range) close to the source node, causing network congestion.

IV. BROADCASTING IN NETWORKS WITH SOURCE QUADRANT INFORMATION

In this section, we study the efficiency of broadcasting in networks with source quadrant information and compute the number of transmissions required to obtain a normalized delay of $\Theta(\frac{1}{M(n)})$. We assume that the nodes have only a local notion of four directions which are not necessarily common to all nodes. That is, the nodes are capable of grouping their neighbors into four different quadrants, where the orientations of the quadrants are chosen independently by different nodes. To model this, we assume that the orientations of the quadrants are uniformly distributed between angles 0 and 2π , and are

chosen independently of the local quadrants at other nodes. We also assume that there is some data embedded in a packet's header that enables an intermediate node to infer coarse geographic source location w.r.t its local quadrants. This could be implemented, for instance, if the packet has the source location embedded in its header and nodes have possibly faulty GPS (see [21]). Thus, the nodes are assumed to have *Source Quadrant Information*. However, we do not assume that the nodes have any state information i.e., they are incapable of remembering any previous transmissions or messages.

To broadcast in networks with limited geographic information, we study broadcasting strategies similar to the schemes presented in Section III. The broadcast strategy follows the random packet forwarding model, but utilizes the location information to direct the packets radially away from the source node, reducing the broadcast redundancy. We again use the multiple independent query model to achieve a normalized delay of $\Theta(\frac{1}{M(n)})$. The broadcast strategy is as follows:

- 1) The source node picks $R(n)$ neighbors uniformly randomly (i.e., $R(n)$ points independently chosen from its neighbor set), and sends the broadcast message to them.
- 2) Each of the nodes, on receiving a request to transmit, retransmit the message and choose exactly one neighbor from the "local" quadrant opposite to the source's quadrant, and request that neighbor to retransmit the message.

A. Broadcast Model with Source Information

Let the source node be at $\mathbf{0}$, and consider any given copy (indexed by $k = 1$ to $R(n)$) of the broadcast message. We denote the transmitting node at the $i-1$ th iteration to be S_{i-1}^k . Since the Source Quadrant Information is available to all nodes, the transmitting node for the i th iteration ($i \geq 2$), $S_i^k = (Z_i^k, \phi_i^k)$ (in polar coordinates) is chosen uniformly from the quadrant opposite to the source. Let us denote the offset angle (from the line joining the source and the node) by θ_i^k and the offset length by X_i^k , as shown in Figure 4(i). Notice that the source quadrant information at the nodes restricts the offset angle to be within $[-\pi/2, \pi/2]$.

The radial progress in the i th jump is defined as the random variable $Y_i^k = \|S_i^k\| - \|S_{i-1}^k\|$, with support in $[0, M(n)]$. As the initial direction of transmission is uniformly distributed over $[0, 2\pi]$, (the source node picks its transmitting nodes uniformly from within its circular radio range), S_i^k are also angularly uniformly distributed (see Theorem 7.1 in Appendix for a formal proof of this claim).

From the geometry of the paths, and due to the availability source quadrant information, we see that the radial progress is always positive in every step. Further, in all steps, we have

$$Y_i^k \geq X_i^k \cos \theta_i^k. \quad (17)$$

See Figure 4(ii) for an illustration of this property. The random variables $X_i^k \cos \theta_i^k$ are i.i.d random variables, with support $[0, M(n)] \subset \mathbb{R}$ and $E(X_i^k \cos \theta_i^k) = dM(n)$, $d > 0$.

Under these conditions, the following theorem provides an upper bound on the number of transmissions required to ensure a delay of $\Theta(\frac{1}{M(n)})$.

Theorem 4.1: Consider $R(n)$ random walks starting from $S_0 = \mathbf{0}$ and any given point $x = (1, \theta^*)$ on the boundary of the compact ball $B_1(0)$. Let $c = \frac{2}{d}$. Then, for $R(n) \geq \frac{1}{M(n)^\alpha}$, $\alpha > 1$,

$$\min_{k \in 1, \dots, R(n), i \in 1, \dots, \frac{c}{M(n)}} \|S_i^k - x\| \leq M(n). \quad (\text{prob.})$$

Proof: We define the hitting time τ^k as the first time-step the path k hits a ball of radius $1 - M(n)$, i.e. the boundary of the set $B_{1-M(n)}(0)$. Notice that the event $\{\tau^k = i\}$ is equivalent to the event that $\{Z_i^k > 1 - M(n) \ \& \ Z_{i-1}^k < 1 - M(n)\}$. By $a \wedge b$, we denote the minimum of the integers a and b .

The probability that no path reaches the point x ,

$$\begin{aligned} & \mathbb{P}\left(\bigcap_{1 \leq k \leq R(n), 1 \leq i \leq \frac{c}{M(n)}} S_i^k \notin B_{M(n)}(x)\right) \\ &= \mathbb{P}\left(\bigcap_{k \in \{1, \dots, R(n)\}} S_{\tau^k \wedge \frac{c}{M(n)}}^k \notin B_{M(n)}(x)\right) \\ &= \left\{ \mathbb{P}\left(S_{\tau^1 \wedge \frac{c}{M(n)}}^1 \notin B_{M(n)}(x)\right) \right\}^{R(n)} \end{aligned} \quad (18)$$

as the paths $1 \leq k \leq R(n)$ are independent and identically distributed.

Consider the probability that the 1st path did not reach the node x in $\frac{c}{M(n)}$ steps.

$$\begin{aligned} & \mathbb{P}\left(S_{\tau^1 \wedge \frac{c}{M(n)}}^1 \notin B_{M(n)}(x)\right) \\ &= \sum_{k=1}^{\infty} \mathbb{P}\left(S_{\tau^1 \wedge \frac{c}{M(n)}}^1 \notin B_{M(n)}(x) \mid \{\tau^1 = k\}\right) \mathbb{P}\left(\tau^1 = k\right) \\ &\leq \sum_{k=1}^{\frac{c}{M(n)}} \mathbb{P}\left(S_k^1 \notin B_{M(n)}(x) \mid \{\tau^1 = k\}\right) \mathbb{P}\left(\tau^1 = k\right) \\ &\quad + \mathbb{P}\left(\tau^1 \geq \frac{c}{M(n)}\right). \end{aligned} \quad (19)$$

To show that the second term of Equation 19, $\mathbb{P}(\tau^1 \geq \frac{c}{M(n)})$ is exponentially decaying (with respect to $\frac{1}{M(n)}$), we first note that (see Equation 17) the radial progress of the first message in $\frac{c}{M(n)}$ steps is lower bounded by $\sum_{i=1}^{\frac{c}{M(n)}} X_i^1 \cos \theta_i^1$. Thus the probability that the radial progress was within the unit circle after $\frac{c}{M(n)}$ steps,

$$\mathbb{P}\left(\tau^1 \geq \frac{c}{M(n)}\right) \leq \mathbb{P}\left(\sum_{i=1}^{\frac{c}{M(n)}} X_i^1 \cos \theta_i^1 \leq 1\right) \leq \exp^{-\gamma \frac{1}{M(n)}}, \quad (20)$$

for some $\gamma > 0$. Equation 20 is by Chernoff's Inequality for sums of random variables $X_i^1 \cos \theta_i^1$, whose mean is $dM(n)$.

Now, given the equivalence of the events $\{\tau^1 = k\}$ and $\{Z_k^1 > 1 - M(n), Z_{k-1}^1 < 1 - M(n)\}$, we demonstrate an upper bound for the first term of Equation 19. We have, for $1 \leq k \leq \frac{c}{M(n)}$,

$$\begin{aligned} & \mathbb{P}\left(S_{\tau^1 \wedge \frac{c}{M(n)}}^1 \notin B_{M(n)}(x) \mid \{\tau^1 = k\}\right) \\ &= \mathbb{P}\left(S_{\tau^1 \wedge \frac{c}{M(n)}}^1 \notin B_{M(n)}(x) \mid \{Z_k^1 > 1 - M(n), Z_{k-1}^1 < 1 - M(n)\}\right) \end{aligned} \quad (21)$$

Since τ^1 is a hitting time, the path must be within an $M(n)$ distance from the boundary. Hence,

$$\begin{aligned} & \mathbb{P}\left(S_k^1 \notin B_{M(n)}(x) \mid \{Z_k^1 > 1 - M(n), Z_{k-1}^1 < 1 - M(n)\}\right) \leq \\ & \mathbb{P}\left(\{\phi_k^1 \in [\phi^* - \frac{M(n)}{4}, \phi^* + \frac{M(n)}{4}]\} \mid \{Z_k^1 > 1 - M(n), Z_{k-1}^1 < 1 - M(n)\}\right) \end{aligned} \quad (22)$$

By the uniform distribution of $\phi_{\tau^1}^1$ (see Theorem 7.1 in Appendix for a proof of this assertion, and for independence of $\phi_{\tau^1}^1$ and τ^1), it follows that

$$\begin{aligned} & \mathbb{P}\left(\{\phi_k^1 \in [\phi^* - \frac{M(n)}{4}, \phi^* + \frac{M(n)}{4}]\} \mid \{Z_k^1 > 1 - M(n), Z_{k-1}^1 < 1 - M(n)\}\right) \\ &= \frac{M(n)}{4\pi}. \end{aligned} \quad (23)$$

By Equations 19,22,23 and for some $k_2 > 0$

$$\mathbb{P}\left(S_{\tau^1 \wedge \frac{c}{M(n)}}^1 \notin B_{M(n)}(x)\right) \leq 1 - k_2 M(n). \quad (24)$$

For $R(n) = \frac{1}{M(n)^\alpha}$, $\forall \alpha > 1$,

$$\begin{aligned} & \mathbb{P}\left(\bigcap_{k \in \{1, \dots, R(n)\}} S_{\tau^k \wedge \frac{c}{M(n)}}^k \notin B_{M(n)}(x)\right) \\ & \leq \{1 - k_2 M(n)\}^{R(n)} \rightarrow 0. \end{aligned} \quad (25)$$

Theorem 4.1 thus follows. \blacksquare

Remark 4.1: Thus, the total number of transmissions $T(n) = (R(n) * \frac{c}{M(n)})$ is less than $(\frac{1}{M(n)})^\alpha$ for any $\alpha > 2$. The results demonstrate that it is sufficient for $(\frac{1}{M(n)})^2$ transmissions to broadcast to any randomly chosen point that is a unit distance away from the source, with local geographic knowledge, even without any suppression of transmissions. However, we note that the broadcast strategy causes a polynomially large (of order $\frac{1}{M(n)}$) number of transmissions around the source node, causing significant congestion, although the congestion is substantially lower, compared to broadcasting with Zero Information, where the number is exponentially large.

V. BROADCASTING WITH LIMITED STATE INFORMATION

In this section, we analyze broadcasting in networks with limited state information. We assume that the nodes in the network are capable of remembering previously received messages and their decision to transmit or to not transmit the received message. However, we assume that the nodes have no knowledge of the position of the neighbors or the source node. In such networks with very little state information, and no location information, we study broadcast strategies that possess broadcast delays of $D(n) = \Theta(\frac{1}{M(n)})$ and compute the number of transmissions required to achieve the order-wise optimal delays. The broadcast scheme we study is a variation of the gossip algorithm presented in [9] where a node decides to retransmit the broadcast message with a probability p , upon the first arrival of the message. The broadcast algorithm we employ is described below.

- 1) In the first iteration, the source node S_0 transmits the message 'm', to all its neighbors, and chooses $C \log \frac{1}{M(n)}$ nodes randomly from its neighbor set, and requests them to retransmit the message.

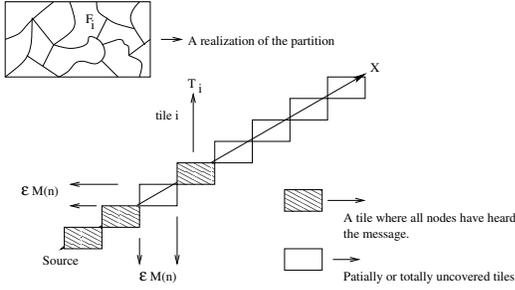


Fig. 5. Branching in Sensor nodes.

- 2) In the next iteration, the chosen nodes transmit their message and choose $C \log \frac{1}{M(n)}$ nodes randomly from their neighbor sets, but nodes that have received the previous broadcast of the message ignore all subsequent broadcasts of the same message. Thus, nodes chosen from regions that had previously heard the message do not transmit.
- 3) The process is repeatedly iterated to spread the query over the network.

Thus, the algorithm employs the state information to suppress redundant transmissions in regions that have previously received the broadcast message. For this “location-less” broadcast scheme, we show that the delay $D(n) = \Theta(\frac{1}{M(n)})$, while the total number of transmissions are $O(\frac{1}{M(n)}^\alpha)$, where $\alpha > 2$.

In the following theorem, we first prove that the broadcast algorithm discussed previously achieves a delay of $\Theta(\frac{1}{M(n)})$. We show this, by choosing any node X , that is a unit distance away from the Source node S_0 and demonstrating that there is a transmission within the radio range of that given node within $\Theta(\frac{1}{M(n)})$ iterations.

By our notations in Section II, we define S_i to be the set of transmitters in iteration i and \mathcal{P}_i to be the set of all transmitters till iteration i .

Theorem 5.1: Let $S_0 = 0$, X be any given point such that $\|X - S_0\| = 1$. Then, for some $\epsilon > 0$, there exists a $0 < C_\epsilon < \infty$, such that for $C = C_\epsilon$,

$$\min_{Y \in \mathcal{P}_{\frac{1}{\epsilon M(n)}}} \|X - Y\| \leq M(n) \quad (\text{prob.}) \quad (26)$$

Proof:

Consider tiles of size $\epsilon M(n) \times \epsilon M(n)$ about the line connecting the source node and X , as in Figure 5. We choose $\epsilon > 0$ such that a transmission (of range $M(n)$) in any tile covers the adjacent tiles as well (it can be seen that for any $\epsilon < \frac{1}{3}$, this condition is satisfied). A tile is defined to be ‘covered’ if all nodes within the tile have received the broadcast message; else it is defined to be ‘uncovered’. Let A_t be the event {Tile T_t covered by time t } and let the event E_t be the event {Some node in Tile T_t was picked as a transmitter}. We require the following lemma.

Lemma 5.1: The probability

$$\mathbb{P}(E_t^c/A_t) = \mathbb{P}(\text{No transmissions in tile } T_k|A_t) \leq M(n) \frac{C\epsilon^2}{\pi}. \quad (27)$$

Proof: Let \mathcal{W} be any partitioning of the tile T_t . Let the partition \mathcal{W} be the union of disjoint sets $F_i, i = 1, \dots, f(n)$,

where the disjoint sets F_i correspond to the incrementally covered regions of the tile T_t , over different transmissions (see Figure 5 for an illustration). Let $l(F_i)$ denote the fraction of the area of F_i in the tile, with $\sum_{i=1}^{f(n)} l(F_i) = 1$. Then,

$$\mathbb{P}(E_t^c|A_t) = \int \mathbb{P}(E_t^c|A_t, \mathcal{W}) d\mu_{A_t}(\mathcal{W}), \quad (28)$$

where $\mu_{A_t}(\mathcal{W})$ is the probability that the partition \mathcal{W} was created by the transmission process. We now derive an uniform upper bound on $\mathbb{P}(E_t^c|A_t, \mathcal{W})$ (which does not depend on \mathcal{W}), and hence, provide an upper bound on L.H.S of (28).

Since we choose $C \log M(n)$ nodes uniformly from an area of $\pi(M(n))^2$, the probability

$$\begin{aligned} \mathbb{P}(E_t^c|A_t, \mathcal{W}) &= \prod_{i=1}^{f(n)} \left(1 - \frac{l(F_i)\epsilon^2}{\pi}\right)^{C \log \frac{1}{M(n)}} \\ &= \prod_{i=1}^{f(n)} M(n)^{\left(C \log \left(\frac{1}{1 - \frac{l(F_i)\epsilon^2}{\pi}}\right)\right)}, \\ &= M(n)^{-C \left(\sum_{i=1}^{f(n)} \log \left(1 - \frac{l(F_i)\epsilon^2}{\pi}\right)\right)}. \end{aligned} \quad (29)$$

As $M(n) < 1$, we now have from (29) $\mathbb{P}(E_t^c|A_t, \mathcal{W}) \leq M(n)^{[C\beta^*]}$, where

$$\begin{aligned} \beta^* &= - \max_{x_i: 1 \leq i \leq f(n)} \sum_{i=1}^{f(n)} g(x_i), \quad \text{s.t.} \quad \sum_{i=1}^{f(n)} x_i = 1, \\ g(x) &= \log \left(1 - (\epsilon^2/\pi)x\right)x, \quad x \in [0, 1], \end{aligned} \quad (30)$$

It can be directly computed to show that $g(x)$ is a negative concave function with $g(0) = 0, g(1) = \log(1 - (\epsilon^2/\pi))$. By using Lagrange Multipliers, it can be shown that for each fixed $f(n)$, the maximum is achieved when $x_i = \frac{1}{f(n)}$, for all i . Thus,

$$\beta^* = - \max_{f(n)} f(n)g\left(\frac{1}{f(n)}\right). \quad (31)$$

Further, we have $\log(1 - (\epsilon^2/\pi)) \leq -(\epsilon^2/\pi)$, and hence, $\beta^* \leq (\epsilon^2/\pi)$. The result now immediately follows. ■

Now, the probability that the tile T_{t+1} was covered by time $t+1$

$$\mathbb{P}(A_{t+1}) \geq \mathbb{P}(A_t \cap E_t), \quad (32)$$

$$= \mathbb{P}(A_t)\mathbb{P}(E_t/A_t), \quad (33)$$

$$= \mathbb{P}(A_t) \left[1 - \mathbb{P}(E_t^c/A_t)\right]. \quad (34)$$

Note that the inequality in (32) is due to the fact that the event $A_t \cap E_t$ implies A_{t+1} , by construction. Utilizing Lemma 5.1 in (34),

$$\begin{aligned} \mathbb{P}(A_{t+1}) &\geq \mathbb{P}(A_t) \left(1 - (M(n))^{C\epsilon^2/\pi}\right) \\ &\geq \left(1 - (M(n))^{C\epsilon^2/\pi}\right)^t \end{aligned} \quad (35)$$

Hence, it follows that

$$\mathbb{P}(A_{\frac{1}{M(n)}}) \geq \left(1 - (M(n))^{C\epsilon^2/\pi}\right)^{\frac{2}{M(n)}} \rightarrow 1, \quad (36)$$

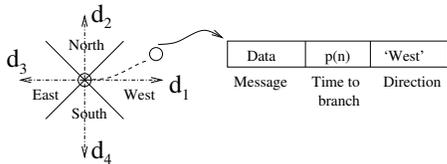


Fig. 6. Branching in Sensor networks with Local Direction Information

for $\frac{C\epsilon^2}{\pi} > 1$. Thus it is seen that by iteration $\frac{K}{M(n)}$, the tile $T_{\frac{K}{M(n)}}$ is covered with high probability. ■

By our construction, we see that in any tile T , the number of transmissions is no greater than $C \log 1/M(n)$. Since the total number of tiles is no greater than $K(\frac{1}{M(n)})^2$, the total number of transmissions $T(n) \leq K_1(\frac{1}{M(n)})^2 \log 1/M(n)$.

Remark 5.1: The results in this section demonstrate that “state information” in the networks can be utilized to simultaneously reduce the number of transmissions, and to distribute the transmissions more uniformly over the network. The proof in this sections show that the state information inherently provides a linear drift, emphasizing the role of suppression in efficient broadcasting. Further, the results can be extended to show that the branching algorithm can spread information uniformly in a two dimensional region. Moreover, uniformly spaced transmissions considerably reduce the congestion in the network.

VI. BROADCASTING OVER A LATTICE WITH LOCAL DIRECTION INFORMATION

In this section, we study the problem of broadcasting to a set of spatially uniform nodes (lattice points) in networks where nodes have no “state” or geographic information, but only a rudimentary sense of local direction. That is, each sensor node in the network has an approximate sense of ‘East’, ‘West’, ‘North’ and ‘South’, formally defined in VI.A. Necessity for such a broadcasting scheme could arise when a spatially uniform sample of an underlying physical process is required by an application at the source node (see figure 3). For example, a sensor network deployed for measuring air quality might require measurements from the sensor network sampled uniformly over the deployed region; and thus, will need to send a query/message to the appropriate subset of nodes. We examine if such queries/messages can be broadcast efficiently with the availability of “local direction” information, and propose a random tree based broadcast protocol that utilizes the local information to spread messages over the network. Under this broadcast scheme, we compute the number of transmissions required to reach a circular advertisement region of radius $A(n)$ about the destination node (a lattice point that is a unit distance away) within a delay of $\Theta(\frac{1}{M(n)})$.

A. Broadcast and Network Model

We assume that sensor nodes in the network have an approximate knowledge about four antipodal directions $d_{(1)}, d_{(j)}, d_{(-1)}, d_{(-j)}$, In particular, the transmitting

nodes have a local estimate of four antipodal directions $\bar{d}_{(1)}, \bar{d}_{(j)}, \bar{d}_{(-1)}, \bar{d}_{(-j)}$, such that for all $l \in \{1, j, -1, -j\}$

$$E(d_{(l)} \cdot \bar{d}_{(l)}) = c, c > 0 \quad \text{and} \quad E(\bar{d}_{(l)}) = d_{(l)}. \quad (37)$$

In other words, we assume that the direction estimates are unbiased and with a positive projection. We note that the expected projection could be differ between directions, however we choose a uniform projection in all directions for notational simplicity.

We also assume that the packet contains information on the direction of travel, and a counter, to keep track of the number of hops traveled by a packet (Figure 6). Without loss of generality, we formally define the four directions to be $d_{(1)} = (1, 0)$, $d_{(j)} = (0, 1)$, $d_{(-1)} = (-1, 0)$ and $d_{(-j)} = (0, -1)$ (See Figure 6). Thus, in a transmission by a node x along the direction d_i , the distance traveled in that transmission is a random variable X , with support $[0, M(n)]d_i \subset \mathbb{R}^2$, and $E(X) = cM(n)d_i, c > 0$.

For networks with local direction information, the randomized tree (branching walk) based broadcast strategy is performed as follows (see Figure 7).

- 1) The source node $S_0 = \mathbf{0}$ transmits a query to a randomly chosen retransmission node in each direction. The packets contain the data, the direction in which they were sent, and the Time to Branch(TTB) counter is set to $p(n)$ (See Figure 6).
- 2) The retransmission nodes check the packet’s TTB counter. If $TTB = 0$, then the retransmission node transmits one query each to the two orthogonal directions to the previous step, and sets $TTB = p(n)$, in the newly created query packets. If $TTB > 0$, then TTB value alone is changed to $TTB - 1$, and the packet is retransmitted along the same direction.

Since the nodes create two queries at every branching, the spatial distribution of the query can be studied as a process indexed by a binary tree. Consider a query sent by the source node along the direction d_i . Let Γ denote an infinite binary tree, where the vertices correspond to the queries generated by repeated branching of the initial query. Let $\Gamma_{(l,k)}$ denote the query at the k th vertex at depth l , with $l \in \mathbb{N}$, and $k \in J_l := \{0, 1, \dots, 2^{l-1} - 1\}$. Let Z_k^l be the position of the query $\Gamma_{(l,k)}$, just before the $i + 1^{th}$ branching. Then,

$$Z_k^l = Z_{\lfloor \frac{k}{2} \rfloor}^{l-1} + Y_k^l, \quad (38)$$

where Y_k^l is the random distance traveled by the query after its l^{th} branching. Hence, the random variable has a support $[0, p(n)M(n)]d_i$ and $E(Y_k^l) = cp(n)M(n)d_i$, where d_i is the direction of travel of the query. As defined in Section II, we denote by \mathcal{S}_i , the set of transmitters in the i th iteration.

Under the model discussed above, we show that the number of transmissions to reach a circular advertisement region of radius $A(n) = M(n)^\alpha, \alpha < \frac{1}{2}$ about any given point $x = (1, \theta^*)$ (in polar coordinates), with a normalized delay of $\Theta(\frac{1}{M(n)})$ is $\frac{1}{M(n)^\gamma}, \forall \gamma > 1$. That is, we show that the number of transmissions $T(n)$ is only marginally greater than an optimal number of transmissions, if an advertising radius

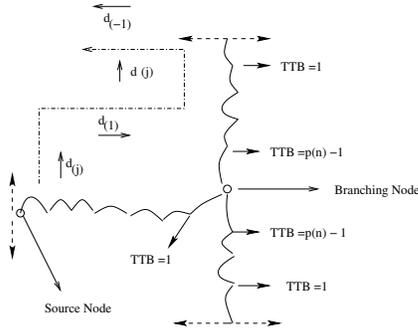


Fig. 7. Illustrates the query branching in sensor networks. Note that the branches do not follow straight lines due to approximate direction knowledge.

of $A(n) = M(n)^\alpha$, $\alpha < \frac{1}{2}$ is allowed. We show this in the following theorem for rational angles.

Theorem 6.1: Consider any point $X = (1, \theta^*)$ on the boundary of a unit ball around the origin. Consider a branching query process as described above. Then, there exists a $0 < b < \infty$ such that $\forall < 1/2$.

$$\min_{Y \in \mathcal{S}_{\frac{b}{M(n)}}} \|Y - X\|_{\mathcal{L}_2} \leq M(n)^\alpha \quad (\text{prob.}) \quad (39)$$

Moreover, by iteration $\frac{b}{M(n)}$, the total number of transmissions

$$T(n) = O\left(\frac{1}{M(n)}^\gamma\right), \forall \gamma > 1. \quad (40)$$

Proof: We first show Theorem 6.1 for $\theta^* \in [0, \frac{\pi}{4}]$ such that $\tan \theta^*$ is rational. The result follows for any $\theta^* \in [0, \frac{\pi}{4}]$ by the density of rationals \mathbb{Q} in \mathbb{R} and by the continuity of $\tan \theta^*$ on $[0, \frac{\pi}{4}]$. For any other $\theta^* \notin [0, \frac{\pi}{4}]$, the result follows, by symmetry.

The main steps of the proof are as follows.

- 1) We employ a $p(n) = \frac{1}{M(n) \log \log \frac{1}{M(n)}}$ to create slowly branching trees,
- 2) We show the existence of a path in the binary tree with a mean angular drift along θ^* .
- 3) We then show that the path lies within a radius $M(n)^\alpha$, $\alpha < \frac{1}{2}$ about the destination X .

Firstly, we describe the construction of the path in the binary tree. Let $\tan \theta^* = \frac{r}{q}$. Recall that the branching occurs exactly once every $p(n) = \frac{1}{M(n) \log \log \frac{1}{M(n)}}$ hops in each query. Further, note that at each branching, exactly two queries are sent along the two perpendicular directions to the original direction along which the query was traveling. That is, if a query traveling along direction $d_{(1)}$ branched, the two new queries would be directed along $d_{(j)}$ and $d_{(-j)}$. Consider the initial queries sent along the direction $d_{(j)}$ and $d_{(-j)}$ by the source node S_0 .

- 1) We denote by $\mathfrak{D}_1 := (d_{(j)}; d_{(1)}; d_{(j)}; d_{(-1)})$, a sequence of the directions of branchings followed by the query, as depicted in Figure 7 (in dotted lines). In particular $(d_{(j)}; d_{(1)}; d_{(j)}; d_{(-1)})$ defines the path of a query through four successive branchings; the direction followed at each branching provided by the sequence of directions. Similarly, we also define another sequence

of branchings $\mathfrak{D}_2 := (d_{(-j)}; d_{(1)}; d_{(j)}; d_{(1)})$. From the construction of the tree, the expected position of the query, after the branchings $(d_{(j)}; d_{(1)}; d_{(j)}; d_{(-1)})$ is $cp(n)M(n)(2d_{(j)})$. The expected position of the query after the sequence of branchings $(d_{(-j)}; d_{(1)}; d_{(j)}; d_{(1)})$ is given by $cp(n)M(n)(2d_{(1)})$.

- 2) Consider the sequence of branchings obtained by following r branchings of type \mathfrak{D}_1 , followed by q branchings of type \mathfrak{D}_2 , i.e., the sequence $\mathfrak{D}_{rq} = (\underbrace{\mathfrak{D}_1; \dots; \mathfrak{D}_1}_{r \text{ terms}}; \underbrace{\mathfrak{D}_2; \dots; \mathfrak{D}_2}_{q \text{ terms}})$. The expected position of the query after the sequence of branchings \mathfrak{D}_{rq} is $cp(n)M(n)(2rd_{(1)} + 2qd_{(j)})$.
- 3) We construct the sequence of branchings formed by following l^* branchings of type branchings \mathfrak{D}_{rq} , where $l^* = \frac{1}{\sqrt{r^2 + q^2} cp(n)M(n)}$. That is, $\mathfrak{D}_{\theta^*} = (\underbrace{\mathfrak{D}_{rq}; \dots; \mathfrak{D}_{rq}}_{l^* \text{ terms}})$.

Note that the expected position of the query after the sequence of branchings \mathfrak{D}_{θ^*} is

$$l^* \times cp(n)M(n)(2rd_{(1)} + 2qd_{(j)}) = (d_{(1)} \cos \theta^* + d_{(j)} \sin \theta^*) \\ = (1, \theta^*) \text{ (in polar coordinates.)}$$

In effect, we construct a path with mean drift along θ^* , by appending a series of branchings. Note that the number of iterations to reach the end of the sequence D_{θ^*} is $l^* \times (r + q) \times 4 \times p(n) = \frac{b}{M(n)}$. Thus, by construction, we show the existence of a path such that the mean position after b iterations is the destination node X . We now show that the position of the path after the sequence of branchings \mathfrak{D}_{θ^*} is within a distance $M(n)^\alpha$ of its mean $X = (1, \theta^*)$, for all $\alpha < .5$, with high probability.

Now, let Γ correspond to a binary tree created by a query along the direction $d_{(j)}$ from the source node. Notice that the position of the path D_{θ^*} is an element of this tree, at depth b . We denote position of the query after the sequence D_{θ^*} by the random variable Z_t^b , where $t \in \{1, \dots, 2^{b-1} - 1\}$. Thus, the position of the query is given by (depth b , leaf t) $Z_t^b = \sum_{i=0}^{b-1} Y_{[\frac{t}{2^i}]}^{b-i}$.

Let $L_1 = \{i : E(Y_{[\frac{t}{2^i}]}^{b-i}) = d_{(1)}\}$, that is, the set of indices such that the query is along direction $d_{(1)}$. Similarly, we define $L_2 = \{i : E(Y_{[\frac{t}{2^i}]}^{b-i}) = d_{(j)}\}$, $L_3 = \{i : E(Y_{[\frac{t}{2^i}]}^{b-i}) = d_{(-1)}\}$ and $L_4 = \{i : E(Y_{[\frac{t}{2^i}]}^{b-i}) = d_{(-j)}\}$.

Since these sets are constructed deterministically, we rewrite sum in (??) as follows.

$$Z_t^b = \sum_{i \in L_1} Y_{[\frac{t}{2^i}]}^{b-i} + \sum_{i \in L_2} Y_{[\frac{t}{2^i}]}^{b-i} + \sum_{i \in L_3} Y_{[\frac{t}{2^i}]}^{b-i} + \sum_{i \in L_4} Y_{[\frac{t}{2^i}]}^{b-i} \quad (41)$$

Notice that $Y_{[\frac{t}{2^i}]}^{b-i}$ for $i \in \{L_r, r = 1 \text{ to } 4\}$ are i.i.d. random variables. For example, $Y_{[\frac{t}{2^i}]}^{b-i}$, $i \in L_1$ is a random variable corresponding to a query along the direction $d_{(1)}$. Thus, each random variable in this set is a sum of $p(n)$ hops along direction $d_{(1)}$. Thus,

$$Y_{[\frac{t}{2^i}]}^{b-i} = d_{(1)} \left(\sum_{m=1}^{p(n)} R_m \right), \quad (42)$$

where R_m are i.i.d. random variables with support $[0, M(n)]$ and mean $cM(n)$. (See discussion in VI.A for the above construction). Since each random variable $Y_{\lfloor \frac{t}{2^r} \rfloor}^{b-i}$ is a sum of $p(n)$ random variables of kind R_m , we have the following claim.

Claim 2: Let $.5 < \beta < 1$. Then,

$$\mathbb{P}\left(\left\|Y_{\lfloor \frac{t}{2^r} \rfloor}^{b-i} - c(M(n)p(n))d_{(1)}\right\| > M(n)(p(n))^\beta\right) \leq e^{-p(n)^{2\beta-1}\epsilon}, \quad (43)$$

for some $\epsilon > 0$.

Proof: By construction,

$$\mathbb{P}\left(\left\|Y_{\lfloor \frac{t}{2^r} \rfloor}^{b-i} - (M(n)p(n)c)d_{(1)}\right\| > M(n)(p(n))^\beta\right) \leq \mathbb{P}\left(\sum_{m=1}^{p(n)} (R_m - M(n)c) > M(n)(p(n))^\beta\right). \quad (44)$$

Let $\tilde{R}_m = \frac{1}{M(n)}R_m$. Then, note that

$$\begin{aligned} & \mathbb{P}\left(\sum_{m=1}^{p(n)} (R_m - M(n)c) > M(n)(p(n))^\beta\right) \\ &= \mathbb{P}\left(\sum_{m=1}^{p(n)} (\tilde{R}_m - c) > (p(n))^\beta\right) \\ &= \mathbb{P}\left(\frac{1}{(p(n))^\beta} \sum_{m=1}^{p(n)} (\tilde{R}_m - c) > 1\right) \leq e^{-p(n)^{2\beta-1}\epsilon}, \quad \epsilon > 0. \quad (45) \end{aligned}$$

A similar inequality can be derived for the negative side as well. We skip the details for brevity. The inequality in (45) follows from the result ([3]) in moderate deviations about the mean, for sums of random variables. ■

Consider the path \mathcal{D}_{θ^*} . It is easily seen that there are $(2q+r)*l^*$ queries in the path along direction $d_{(1)}$, $(2r+q)*l^*$ in the path along direction $d_{(j)}$, $(r)*l^*$ in the path along direction $d_{(-1)}$ and $(q)*l^*$ in the path along direction $d_{(-j)}$. Note that this implies that for the first term on the R.H.S of (41)

$$\mathbb{P}\left(\left\|\sum_{i \in L_1} Y_{\lfloor \frac{t}{2^r} \rfloor}^{b-i} - \frac{2q+r}{\sqrt{r^2+q^2}}d_{(1)}\right\| > K(p(n))^{\beta-1}\right) \leq e^{-p(n)^{2\beta-1}\epsilon_1}, \quad (46)$$

for some $\epsilon_1 > 0$, and $K < \infty$. Using a similar bound for all the terms on the R.H.S of (41), and noting that

$$\begin{aligned} X &= \frac{2q+r}{\sqrt{r^2+q^2}}d_{(1)} + \frac{2r+q}{\sqrt{r^2+q^2}}d_{(j)} \\ &+ \frac{r}{\sqrt{r^2+q^2}}d_{(-1)} + \frac{q}{\sqrt{r^2+q^2}}d_{(-j)}, \quad (47) \end{aligned}$$

we find that

$$\mathbb{P}\left(\|Z_t^b - X\| > K_1(p(n))^{\beta-1}\right) \leq e^{-p(n)^{2\beta-1}\epsilon_2} \quad (48)$$

for some $\epsilon_2 > 0$, and $K_1 < \infty$. Since $p(n) = \frac{1}{M(n) \log \log \frac{1}{M(n)}}$, the quantity $K(p(n))^{\beta-1} = O(M(n)^\alpha)$ for all $\alpha < 1 - \beta$, and thus, (39) follows.

The total number of transmissions in any binary tree by iteration $\frac{b}{M(n)}$ is given by $p(n) \times 2^{K \log \log \frac{1}{M(n)}}$, where

M(n)	Sub-Critical		Super-Critical	
	Parameter	Prob.	Parameter	Prob.
0.11	c = 1.4	0.12	c = 2.0	0.99
0.09	c = 1.4	0.09	c = 1.9	0.93
0.07	c = 1.4	0.02	c = 2.0	0.99

TABLE II

ZERO INFORMATION - SUCCESS PROBABILITY WITH $15/M(n)$ ITERATIONS

M(n)	Sub-Critical		Super-Critical	
	Parameter	Prob.	Parameter	Prob.
0.11	$\gamma = 1.5$	0.25	$\gamma = 3.0$	1.00
0.09	$\gamma = 1.5$	0.2	$\gamma = 2.7$	0.97
0.07	$\gamma = 1.5$	0.07	$\gamma = 2.7$	0.97

TABLE III

SOURCE QUADRANT INFORMATION - SUCCESS PROBABILITY WITH $2/M(n)$ ITERATIONS

$K \log \log \frac{1}{M(n)}$ is the depth of the binary tree. Notice that we create four binary trees, and hence the total number of transmissions $T(n) = 4p(n) * \log \frac{1}{M(n)}^K$, which is order-wise smaller than $\frac{1}{M(n)}^\gamma$ for all $\gamma > 1$. ■

Remark 6.1: Thus, the results in this section show that even with approximate local direction information, the number of transmissions to reach an advertisement region of $\sqrt{M(n)}$ is only $\Theta(\frac{1}{M(n)}^\gamma)$. That is, a polynomial number of transmissions are sufficient to spread queries efficiently to lattice points in networks with approximate local direction.

VII. SIMULATION RESULTS

In this section, we provide simulation results for the strategies considered in this paper. In all the simulations, we set the source location to be at $(0, 0)$. For the first three broadcast strategies, the destination is chosen to be at $(1, 0)$. For spatial sampling (broadcasting on a lattice), we choose

M(n)	Sub-Critical		Super-Critical	
	Parameter	Prob.	Parameter	Prob.
0.11	$C = 1$	0.26	$C = 2.0$	0.99
0.09	$C = 1$	0.25	$C = 2.0$	1.00
0.07	$C = 1$	0.12	$C = 2.0$	1.00

TABLE IV

STATE INFORMATION - SUCCESS PROBABILITY WITH $5/M(n)$ ITERATIONS

M(n)	Sub-Critical		Super-Critical	
	Parameter	Prob.	Parameter	Prob.
0.06	$\alpha = 0.8$	0.18	$\alpha c = 0.4$	0.94
0.04	$\alpha = 0.8$	0.16	$\alpha c = 0.4$	0.90
0.02	$\alpha = 0.8$	0.20	$\alpha c = 0.4$	0.98

TABLE V

SPATIAL SAMPLING BY BRANCHING - SUCCESS PROBABILITY WITH $15/M(n)$ ITERATIONS

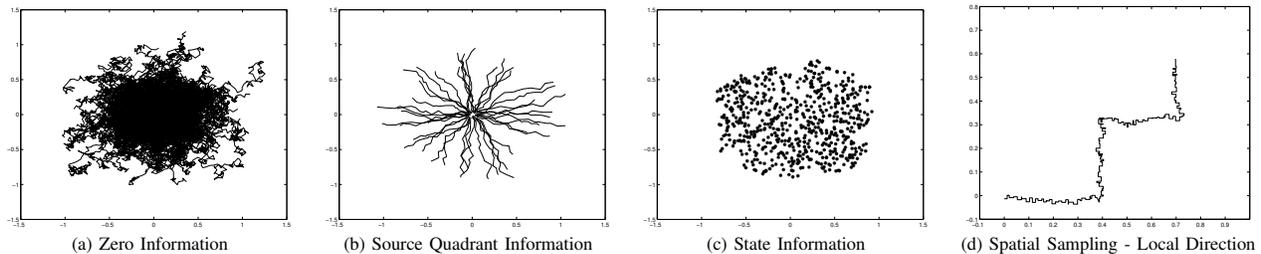


Fig. 8. Sample Paths of Broadcasts in Networks with Local Information

the destination to be at $(.7, .7)$ (for better representation). For each of the strategies, we provide simulation results to show the probability of “success” (appropriately defined for each strategy) for varying parameters and averaged over 50 runs. The transmission radius is chosen such that the number of hops between the source and destination is about 10 – 15.

In Table II, we have provided the probability that a query reaches within an $M(n)$ distance of the destination (success) within $\Theta(1/M(n))$ for the case where there is no information. We have earlier shown that an exponential number of queries are necessary and sufficient for broadcasting without information. To illustrate this by simulation, we have chosen two constants $c_i, i = 1, 2$ and the number of parallel queries sent by the source is $c_i^{1/M(n)}$. The table shows that if c_1 is chosen small enough (but still resulting in an exponential number of queries), the probability of success is small, while a larger value of c_2 results in a success probability that is close to ‘1’, as predicted in Section III. A sample path of the parallel query strategy is illustrated in Figure 8.

In Table III, nodes have source-quadrant information, thus requiring only a polynomial number of parallel queries (with the exponent being 2). In the table we have chosen two growth exponents $\gamma_1 < 2 < \gamma_2$ (i.e., the number of parallel queries is $(1/M(n)^\gamma)$), and the results demonstrate a “sub-critical” rate and a “super-critical” rate (i.e., the probabilities are close to ‘0’ or ‘1’ respectively). In Table IV, a similar result has been plotted for the suppression based strategy (local state-information), with up to $C \log(1/M(n))$ new transmitters chosen (prior to suppression). Again, we can see the sub-critical and super-critical behavior. Finally, in Table V, we have shown a sub-critical and super-critical behavior for lattice flooding, with an advertisement radius $A(n) = M(n)^\alpha$. We have chosen $\alpha_1 < 0.5 < \alpha_2$ to show that the advertisement region needs to be large enough for success. Sample paths of all the strategies described above are illustrated in Figure 8.

APPENDIX

Consider the spatial position (in polar coordinates) $S_i^k = (Z_i^k, \phi_i^k)$ of the k^{th} packet after i steps, as described in Section IV. Under the model described in that section, we show here that the radial progress of a packet is independent of its angular position, and the angular position at the moment it reaches the unit ball is uniform in $[0, 2\pi]$.

Theorem 7.1: Let (Z_i, ϕ_i) be the polar coordinates of the k^{th} packet after i steps, for any k . Then, Z_i is independent of $\phi_i \forall i$. Also, ϕ_τ is uniformly distributed in $[0, 2\pi]$.

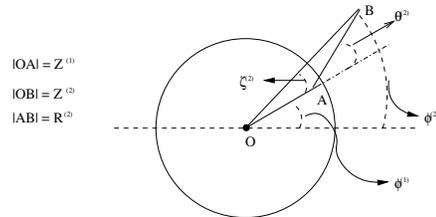


Fig. 9. Independence of the angle and radial progress.

Proof: We first prove that Z_2 is independent of ϕ_2 , and show that ϕ_2 is uniform. We then extend this argument to show the result for any Z_i and ϕ_i .

Consider Figure 9. Initially, the packet chooses a random neighbor $A = (Z_1, \phi_1)$ to retransmit. Notice that as the point is uniformly chosen from within the whole circle, Z_1 is independent of ϕ_1 , and ϕ_1 is uniform in $[0, 2\pi]$. In the next hop, it chooses a neighbor B , from a local quadrant of A , the quadrant opposite to the source O . Since the orientations of the local quadrants are uniformly distributed in $[0, 2\pi]$, and the point B is chosen uniformly from the quadrant opposite to the source, we have the angle $\theta_{(2)}$ to distributed (not uniformly) between $[-\pi/2, \pi/2]$. More importantly, we notice that $\theta_{(2)}$ is independent of both Z_1 and ϕ_1 , i.e., it is independent of the position of A . Now,

$$\zeta^{(2)} = \tan^{-1} \frac{R^{(2)} \sin \theta^{(2)}}{Z_1 + R^{(2)} \cos \theta^{(2)}}, \quad (49)$$

$$Z_2 = \sqrt{(Z_1 + R^{(2)} \cos \theta^{(2)})^2 + (R^{(2)} \sin \theta^{(2)})^2}. \quad (50)$$

Notice that as $\zeta^{(2)}$ is a function of variables that are independent of ϕ_1 , $\zeta^{(2)}$ is independent of ϕ_1 . This implies that

$$\phi_2 = [\phi_1 + \zeta^{(2)}] \text{ mod } 2\pi \quad (51)$$

is uniform in $[0, 2\pi]$ irrespective of the distribution of $\zeta^{(2)}$.

In order to show that Z_2 is independent of ϕ_2 , we show that the distribution of the random variable $f_{\phi_2|Z_2}$ is also uniform in $[0, 2\pi]$. Now, since ϕ_1 is independent of both Z_2 and $\zeta^{(2)}$, we see that

$$\begin{aligned} f_{\phi_2|Z_2} &= f_{[\phi_1 + \zeta_{(2)}] \text{ mod } 2\pi | Z_2} \\ &= f_{\phi_1|Z_2} * f_{\zeta_{(2)}|Z_2, \phi_1} \\ &= f_{\phi_1} * f_{\zeta_{(2)}|Z_2} \end{aligned} \quad (52)$$

where “ $*$ ” implies a wrap-around convolution over the domain $[0, 2\pi]$. We define this formally as follows.

$$(f * g)(t) = \int_0^{2\pi} f(\tau)g(\{t - \tau\} \text{ mod } 2\pi) d\tau. \quad (53)$$

Again, as ϕ_1 is uniformly distributed in $[0, 2\pi]$, and independent of other random variables, it follows that $f_{\phi_2|Z_2}$ is uniformly distributed, thus implying that ϕ_2 is indeed independent of Z_2 .

For any step i , the proof is similar. We construct a $\zeta_{(i)}$ such that $\phi_i = \phi_1 + \zeta_{(i)}$. By means of a similar argument, we show that as $f_{\phi_i|Z_i}$ is uniformly distributed and equal to f_{ϕ_i} , utilizing the fact that f_{ϕ_1} is independent of other quantities, and is also uniformly distributed.

To show that ϕ_τ is uniform in $[0, 2\pi]$, notice that the event $\tau = k$ is equivalent to the event $Z_{k-1} \leq 1 - M(n), Z_k \geq 1 - M(n)$. Also, as ϕ_1 is independent of $\zeta_{(k)}$ and (Z_k, Z_{k-1}) ,

$$\begin{aligned} f_{\phi_k|Z_k, Z_{k-1}} &= f_{[\phi_1 + \zeta_{(k)}] \bmod 2\pi | Z_k, Z_{k-1}} \\ &= f_{\phi_1 | Z_k, Z_{k-1}} * f_{\zeta_{(k)} | Z_k, Z_{k-1} \phi_1} \\ &= f_{\phi_1} * f_{\zeta_{(k)} | Z_k, Z_{k-1}}. \end{aligned} \quad (54)$$

By Equation 51, this implies that the distribution of ϕ_k^1 given that $\tau = k$ is uniform in $[0, 2\pi]$ for all k , and hence it follows that ϕ_τ is uniform in $[0, 2\pi]$. ■

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